

# General Certificate of Education (A-level) June 2011 

Mathematics
MS03

## (Specification 6360)

## Statistics 3

## Final

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## Key to mark scheme abbreviations

| M | mark is for method |
| :--- | :--- |
| m or dM | mark is dependent on one or more M marks and is for method |
| A | mark is dependent on M or m marks and is for accuracy |
| B | mark is independent of M or m marks and is for method and accuracy |
| E | mark is for explanation |
| Jor ft or F | follow through from previous incorrect result |
| CAO | correct answer only |
| CSO | correct solution only |
| AWFW | anything which falls within |
| AWRT | anything which rounds to |
| ACF | any correct form |
| AG | answer given |
| SC | special case |
| OE | or equivalent |
| A2,1 | 2 or 1 (or 0) accuracy marks |
| $-x$ EE | deduct $x$ marks for each error |
| NMS | no method shown |
| PI | possibly implied <br> SCA |
| substantially correct approach |  |
| cf | candidate |
| dp | significant figure(s) |
| decimal place(s) |  |

## No Method Shown

Where the question specifically requires a particular method to be used, we must usually see evidence of use of this method for any marks to be awarded.

Where the answer can be reasonably obtained without showing working and it is very unlikely that the correct answer can be obtained by using an incorrect method, we must award full marks. However, the obvious penalty to candidates showing no working is that incorrect answers, however close, earn no marks.

Where a question asks the candidate to state or write down a result, no method need be shown for full marks.
Where the permitted calculator has functions which reasonably allow the solution of the question directly, the correct answer without working earns full marks, unless it is given to less than the degree of accuracy accepted in the mark scheme, when it gains no marks.

Otherwise we require evidence of a correct method for any marks to be awarded.

| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 1 (a) | $\begin{aligned} & \mathrm{H}_{0}: p=0.25(25 \%) \\ & \mathrm{H}_{1}: p>0.25 \\ & \mathrm{SL} \quad \alpha=0.02(2 \%) \end{aligned}$ | B1 |  | Both |
|  | $\text { CV } \quad z=2.05 \text { to } 2.06$ | B1 |  | AWFW Allow 2.32 to 2.33 if $\mathrm{H}_{1}: p \neq 0.25$ |
|  | $\hat{p}=\frac{108}{375}=\mathbf{0 . 2 8 8}$ | B1 |  | CAO |
|  | $z=\frac{0.288-0.25}{\sqrt{\frac{0.25 \times 0.75}{375}}}=\mathbf{1 . 7 0}$ | $\begin{gathered} \text { M1 } \\ \text { A1 } \end{gathered}$ |  | Allow use of 0.288 in denominator AWRT |
|  | or |  |  | $\mathrm{P}(X \geq 108 \mid n=375, p=0.25)=\mathbf{0 . 0 5 2}$ |
|  | $z=\frac{108(-0.5)-93.75}{\sqrt{375 \times 0.25 \times 0.75}}=\mathbf{1 . 7 0}(\text { or } 1.64)$ | $\begin{gathered} \text { (M1) } \\ \text { (A1) } \end{gathered}$ |  | Allow use of 0.288 in denominator AWRT |
|  | Thus, no evidence, at $2 \%$ level, to support consumer report's claim | AF1 | 6 | F on CV and $z$-value or F on $2 \%$ and probability |
| (b) | Can be considered to be a random sample | B1 | 1 |  |
|  |  | Total | 7 |  |


| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 2 (a) | $98 \% \Rightarrow z=2.32$ to 2.33 | B1 |  | AWFW (2.3263) |
|  | CI for $\lambda$ is: $\hat{\lambda} \pm z \times \sqrt{\hat{\lambda}} \quad \text { or } \quad \bar{x} \pm z \times \sqrt{\frac{\bar{x}}{n}}$ ie | M1 |  | Form; allow $\hat{\lambda} \pm z \times \sqrt{\frac{\hat{\lambda}}{n}}$ |
|  | $108 \pm 2.3263 \times \sqrt{108}$ <br> or | AF1 |  | F on $z$ only; allow $108 \pm z \times \sqrt{\frac{108}{13}}$ |
|  | $\frac{108}{13} \pm 2.3263 \times \sqrt{\frac{108}{13^{2}}}$ | (AF1) |  | F on $z$ only; allow $\frac{108}{13} \pm z \times \sqrt{\frac{108}{13}}$ |
|  | Dividing by 13 or equivalent to obtain a correct numerical expression | A1 |  | May be implied |
|  | Thus $8.31 \pm 1.86$ or $(6.45,10.2)$ | A1 | 5 | AWRT |
|  | Note: <br> For incorrect numerical expressions the maximum marks are B1 M1 AF1 A0 A0 |  |  |  |
| (b) | 1 per 24 hours $\Rightarrow 7$ per week |  |  |  |
|  | CI includes 7 | BF1 |  | F on (a); must use 7 or $1 \mathrm{v} \mathrm{CI} / 7$ |
|  | No reason, at $2 \%$ level, to dispute station officer's claim | Bdep 1 | 2 | Or equivalent Dependent on BF1 |
|  |  | Total | 7 |  |

MS03 (cont)

| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 3 (a)(i)(ii) | $\mathrm{P}(\mathrm{G})=\mathbf{0 . 1 5}$ | B1 | 1 | CAO |
|  | $\mathrm{P}(\mathrm{A} \cap \leq 1)=0.60 \times 0.55=0.33$ | B1 | 1 | CAO |
| (iii) | $\begin{aligned} \mathrm{P}(\leq 24)= & (0.60 \times 0.80)+(0.25 \times 0.85) \\ & +(0.15 \times 0.75) \end{aligned}$ | M1 |  | May be implied |
|  | $=0.48+0.2125+0.1125=\mathbf{0 . 8 0 5}$ | A1 | 2 | CAO |
| (iv) | $\mathrm{P}(\mathrm{~B} \mid \leq 24)=\frac{\mathrm{P}(\mathrm{~B} \cap \leq 24)}{\mathrm{P}(\leq 24)}$ | M1 |  | Used; may be implied |
|  | $=\frac{0.25 \times 0.85}{(\mathrm{iii})}=\frac{0.2125}{0.805}$ | AF1 |  | F on (iii) |
|  | $=0.264$ | A1 | 3 | AWRT |
| (b)(i) | $\mathrm{P}(3$ @ $\mathrm{B} \mid \leq 24)=[(\mathrm{a})(\mathrm{iv})]^{3}$ | M1 |  | Used; may be implied |
|  | $=0.018$ to 0.0185 | A1 | 2 | AWFW (0.01839) |
| (ii) | $\begin{aligned} & \mathrm{P}(\text { same station } \mid \leq 24) \\ & =[\mathrm{P}(\mathrm{~A} \mid \leq 24)]^{3}+(\mathrm{b})(\mathrm{i})+[\mathrm{P}(\mathrm{G} \mid \leq 24)]^{3} \end{aligned}$ | M1 |  | Used; may be implied |
|  | $=\left(\frac{0.48}{0.805}\right)^{3}+(\mathrm{b})(\mathrm{i})+\left(\frac{0.1125}{0.805}\right)^{3}$ | M1 M1 |  | At least 1 term correct; allow (b)(i) providing it is a (cond prob) ${ }^{3}$ <br> All 3 terms correct |
|  | $=0.2120+0.0184+0.0027=\mathbf{0 . 2 3 3}$ | A1 | 4 | AWRT (0.23312) |
|  |  | Total | 13 |  |

MS03 (cont)

| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 4 | 95\% $\Rightarrow z=1.96$ | B1 |  | CAO (AWRT from calculator) |
|  | $\text { Require } \quad 2 \times \frac{1.96 \sigma}{\sqrt{n}} \leq 0.2 \mu$ | M1 |  | Used; may be implied <br> Allow 'no $2 \times$ ' <br> Allow ' $=$ sign' throughout |
|  | Thus $\quad 2 \times \frac{1.96}{\sqrt{n}} \times \frac{\mu}{2} \leq 0.2 \mu$ | M1 |  | Use of $\sigma=\frac{\mu}{2}$; may be implied Allow 'no $2 \times$ ' |
|  | Thus $\quad \sqrt{n} \geq \frac{1.96}{2}$ | M1 |  | Attempt at solution for $\sqrt{n}$ or $n$ |
|  | Thus $n \geq 96.04$ |  |  |  |
|  | Thus, to nearest 10; $n=\mathbf{1 0 0}$ | A1 | 5 | CAO |
|  |  | Total | 5 |  |

MS03 (cont)

| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 5 | E-mails are selected: randomly independently | $\begin{aligned} & \text { B1 } \\ & \text { B1 } \end{aligned}$ |  |  |
|  | $99 \% \Rightarrow z=2.57$ to 2.58 | B1 |  | AWFW (2.5758) |
|  | $\hat{p}_{G}=\frac{72}{160}=\mathbf{0 . 4 5} \quad \hat{p}_{H}=\frac{102}{250}=\mathbf{0 . 4 0 8}$ | B1 |  | CAO both; ignore notation |
|  | Approximate CI for $p_{G}-p_{H}$ is: |  |  |  |
|  | $\left(\hat{p}_{G}-\hat{p}_{H}\right) \pm z \times \sqrt{\frac{\hat{p}_{G}\left(1-\hat{p}_{G}\right)}{n_{G}}+\frac{\hat{p}_{H}\left(1-\hat{p}_{H}\right)}{n_{H}}}$ | M1 m1 |  | Form used <br> Standard deviation term |
|  | Thus: |  |  |  |
|  | $(0.45-0.408) \pm$ |  |  |  |
|  | $1.96 \times \sqrt{\frac{0.45 \times 0.55}{160}+\frac{0.408 \times 0.592}{250}}$ | AF1 |  | Or equivalent <br> F on $\hat{p}_{G}, \hat{p}_{H}$ and $z$ |
|  | Thus: |  |  |  |
|  | $\begin{aligned} & 0.042 \pm 0.129 \\ & \text { or } \\ & (-0.09,0.17) \end{aligned}$ | A1 | 8 | CAO/AWRT <br> or <br> AWRT |
|  | Note: <br> If a pooled estimate of variance is used, then the maximum marks are B1 B1 B1 B1 M1 m0 AF0 A0 (5) |  |  |  |
|  |  | Total | 8 |  |



MS03 (cont)


MS03 (cont)

| Q | Solution | Marks | Total | Comments |
| :---: | :---: | :---: | :---: | :---: |
| 8(a)(i) | $\begin{aligned} & \mathrm{H}_{0}: \mu_{\mathrm{A}}=\mu_{\mathrm{B}} \\ & \mathrm{H}_{1}: \mu_{\mathrm{A}} \neq \mu_{B} \end{aligned}$ | B1 |  | Both; allow suffices of $1 \& 2$ or $X \& Y$ |
|  | $\begin{array}{ll} \mathrm{SL} & \alpha=0.05(5 \%) \\ \mathrm{CV} & z= \pm \mathbf{1 . 9 6} \end{array}$ | B1 |  | CAO (AWRT from calculator) <br> Allow (+)1.96 |
|  | $z=\frac{\bar{x}-\bar{y}}{\sigma \times \sqrt{\frac{1}{n_{A}}+\frac{1}{n_{B}}}}$ | M1 |  | Attempted use; or equivalent Allow $\sigma_{A}$ and $\sigma_{B}$ |
|  | $z=\frac{3770-3695}{285 \times \sqrt{\frac{1}{80}+\frac{1}{120}}}$ | A1 |  |  |
|  | $=1.82$ | A1 |  | AWRT (1.82321) |
|  | No evidence, at $5 \%$ level, to suggest that there is a difference | AF1 | 6 | F on CV and z-value |
| (ii) | Large samples (so CLT is applicable) | B1 | 1 |  |
| (b)(i) | $z=\frac{\bar{x}-\bar{y}}{285 \times \sqrt{\frac{1}{80}+\frac{1}{120}}}= \pm 1.96$ | M1 |  | Equating z-term to 1.96 |
|  | $\text { Thus } \quad \begin{aligned} (\bar{x}-\bar{y}) & = \pm 1.96 \times 41.13616 \\ & = \pm \mathbf{8 0 . 6 3} \end{aligned}$ | A1 | 2 | Requires a convincing deduction AG |
| (ii) | $\mathrm{P}($ Type II error $)=\mathrm{P}\left(\right.$ accept $\mathrm{H}_{0} \mid \mathrm{H}_{0}$ false $)=$ | B1 |  | Used or stated; may be implied |
|  | $\mathrm{P}\left(-80.63<(\bar{x}-\bar{y})<80.63 \mid \mu_{A}-\mu_{B}=125\right)$ |  |  | Accept $(\bar{x}-\bar{y})<80.63$ |
|  | $\mathrm{P}\left(Z<\frac{80.63-125}{285 \times \sqrt{\frac{1}{80}+\frac{1}{120}}}\right)$ | M1 |  | $\begin{aligned} & -80.63 \Rightarrow z=-5.00 \\ & \Rightarrow \text { probability } \approx 0 \end{aligned}$ |
|  | $=\mathrm{P}(\mathrm{Z}<-\mathbf{1 . 0 8})$ | A1 |  | AWRT; ignore sign |
|  | = 0.14 | A1 | 4 | AWRT (0.14038) |
|  |  | Total | 13 |  |
|  | TOTAL |  | 75 |  |

